

195a.

BYLAAG A.

Indien

$$(2.40') \quad \lim_{N \rightarrow \infty} N^{-1} \sum_h (a_h - \bar{\Psi}_N)^2 \geq \varepsilon > 0 \quad \text{en}$$

$$(2.42') \quad N^{-1} \sum_h (\Psi_{Nh} - \bar{\Psi}_N)^4 = o(N)$$

dan voldoen die variante \hat{t}_i aan die voorwaarde

$$(2.50) \quad \lim_{N \rightarrow \infty} E |\hat{t}_i|^{2+\delta} < \infty \quad \text{vir } i=1,2,\dots,k.$$

Bewys: Stel $M_N = \sum_h (a_h - \bar{\Psi}_N)^2$ en

$$P_N = \sum_h (a_h - \bar{\Psi}_N)^4 \leq \sum_h (\Psi_{Nh} - \bar{\Psi}_N)^4.$$

Op analoë wyse as in lemma 2.2 volg dat

$$E(\underline{t}_i - E\underline{t}_i)^4 = \frac{n_i(N-n_i)(N^2+N-6Nn_i+6n_i^2)}{N(N-1)(N-2)(N-3)} \cdot P_N + \frac{3n_i(n_i-1)(N-n_i)(N-n_i-1)}{N(N-1)(N-2)(N-3)} \cdot M_N^2.$$

Uit (2.4) volg

$$\begin{aligned} \hat{t}_i &= (N-n_i)^{\frac{1}{2}} [\underline{t}_i - E\underline{t}_i] [n_i(N-n_i)N^{-1}(N-1)^{-1}M_N]^{-\frac{1}{2}} \\ &= [\underline{t}_i - E\underline{t}_i] [n_i(N-1)^{-1}M_N]^{-\frac{1}{2}}. \end{aligned}$$

$$\begin{aligned} E |\hat{t}_i|^4 &= E [\underline{t}_i - E\underline{t}_i]^4 [n_i(N-1)^{-1}M_N]^{-2} \\ &= \frac{(N-1)(N-n_i)(N^2+N-6Nn_i+6n_i^2)}{n_iN(N-2)(N-3)M_N^2} \cdot P_N + \\ &\quad + \frac{3(N-1)(N-n_i)(N-n_i-1)(n_i-1)}{n_iN(N-2)(N-3)} \\ &= o(1) \end{aligned}$$

$$\begin{aligned} \text{want } P_N M_N^{-2} &\leq N^{-1} \sum_h (\Psi_{Nh} - \bar{\Psi}_N)^4 N^{-1} [N^{-1} \sum_h (a_h - \bar{\Psi}_N)^2]^{-2} \\ &= o(1) \text{ uit (2.40') en (2.42')}. \end{aligned}$$

Gevolgtlik geld $E |\hat{t}_i|^{2+\delta} = o(1)$ met $0 < \delta \leq 2$.

Opmerking.

Onder voorwaarde (2.40) geld voorwaarde (2.40') met waarskynlikheid een.

195b.

BYLAAG B.

Onder die voorwaardes (3.10) - (3.13), (3.28) en

$$(3.14') \quad \begin{aligned} |J(H)| &\leq K[H(1-H)]^{-\frac{1}{4}+\delta} \\ |J^{(m)}(H)| &\leq K[H(1-H)]^{-m-\frac{1}{2}+\delta}, \quad m=1,2 \quad \text{en } \delta > 0 \end{aligned}$$

geld die voorwaarde

$$(3.76) \quad \lim_{N \rightarrow \infty} E|\hat{t}_i|^{2+\delta} < \infty \quad \text{vir } i=1,2,\dots,k \quad \text{en } 0 < \delta \leq 2.$$

Bewys:
$$E|\hat{t}_i|^{2+\delta} = \left[\frac{N-n_i}{N\text{var}(t_i)} \right]^{1+\frac{1}{2}\delta} E|t_i - Et_i|^{2+\delta} \quad \text{uit (3.56).}$$

Aangesien $0 < \lim_{N \rightarrow \infty} N\text{var}(t_i) < \infty$ uit (3.68), is dit

voldoende om te bewys dat $N^{1+\frac{1}{2}\delta} E|t_i - Et_i|^{2+\delta} = O(1)$.

Uit §3.4 opmerking 3, (3.60) en (3.62) volg:

$$E(t_i) = A^{(i)} + o(N^{-\frac{1}{2}}) \quad \text{met } A^{(i)} \text{ gedefinieer}$$

deur (3.61). Nou is

$$\begin{aligned} N^{1+\frac{1}{2}\delta} E|t_i - Et_i|^{2+\delta} &= N^{1+\frac{1}{2}\delta} E|B_{1N}^{(i)} + B_{2N}^{(i)} + o_p(N^{-\frac{1}{2}})|^{2+\delta} \\ &= N^{1+\frac{1}{2}\delta} E|\sum_g N_g [n_i^{-1} \sum_{j=1}^{n_i} \tilde{B}_g(x_{ij}) - n_g^{-1} \sum_{j=1}^{n_g} \tilde{B}_g(x_{gj})] + o_p(N^{-\frac{1}{2}})|^{2+\delta}. \end{aligned}$$

Die resultaat volg m.b.v. lemma 3.11 saam met

(3.1) en (3.3) indien bewys kan word dat

$$E|n_i^{-\frac{1}{2}} \sum_{j=1}^{n_i} \tilde{B}_g(x_{ij})|^{2+\delta} = O(1).$$

Onder voorwaarde (3.14') word hieraan voldoen

vir $\delta = 2$, want:

$$\begin{aligned} E|n_i^{-\frac{1}{2}} \sum_j \tilde{B}_g(x_{ij})|^4 &= E[n_i^{-2} \sum_j \sum_{j'} \sum_{j''} \sum_{j'''} \tilde{B}_g(x_{ij}) \tilde{B}_g(x_{ij'}) \tilde{B}_g(x_{ij''}) \tilde{B}_g(x_{ij'''})] \quad \text{waar} \\ &\quad \text{alle } j \text{'s die waardes } 1,2,\dots,n_i \text{ deurloop} \\ &= n_i^{-2} [\sum_{j,j'} E\tilde{B}_g^2(x_{ij}) \tilde{B}_g^2(x_{ij'}) + \sum_j E\tilde{B}_g^4(x_{ij})] \quad \text{want } x_{ij} \text{ en} \\ &\quad x_{ij'} \text{ is onderling onafhanklik vir } j \neq j' \text{ en } E\tilde{B}_g(x_{ij}) = 0 \\ &= O(1) \end{aligned}$$

aangesien:

$$n_i^{-2} \sum_j \sum_{j'} E \tilde{B}_g^2(x_{ij}) E \tilde{B}_g^2(x_{ij'}) = O(1) \text{ m.b.v. (3.27) en}$$

onder op bladsy 72, en

$$n_i^{-2} \sum_j E \tilde{B}_g^4(x_{ij}) = O(1) \text{ soos blyk uit die volgende:}$$

$$\begin{aligned} |B_g(x_{ij})| &= \left| \int_{x_0}^{x_{ij}} J' [H(x)] dF_g(x) \right| \\ &\leq K \left| \int_{x_0}^{x_{ij}} J' [H(x)] dH(x) \right| \\ &\leq K [|J[H(x_{ij})]| + |J[H(x_0)]|] \\ &\leq K' [H(x_{ij}) \{1-H(x_{ij})\}]^{\delta - \frac{1}{4}} + O(1) \text{ uit (3.14')}. \\ E[B_g(x_{ij})]^4 &\leq K'' \int_0^1 [H(1-H)]^{4(\delta - \frac{1}{4})} dH + O(1) \\ &= K'' \int_0^1 [H(1-H)]^{-1+4\delta} dH + O(1) \\ &= O(1), \text{ sodat} \end{aligned}$$

$$E[\tilde{B}_g(x_{ij})]^4 = O(1).$$

Hiermee is die bewys voltooi.

Opmerkings.

1). Waarskynlik is die meer beperkende voorwaarde (3.14') in vergelyking met (3.14) onnodig. Let in hierdie verband op dat $\lim_{N \rightarrow \infty} E \tilde{B}_g^2(x_{ij}) < \infty$ onder (3.14) - sien C+S(1958). Verder is nodig vir bogenoemde afleiding dat $\lim_{N \rightarrow \infty} n_i^{-2} \sum_j E \tilde{B}_g^4(x_{ij}) < \infty$ waarby reeds

$$\lim_{N \rightarrow \infty} n_i^{-1} \sum_j E \tilde{B}_g^4(x_{ij}) < \infty \text{ as (3.14') geld.}$$

Daar kan ook op gewys word dat

$$(3.13) \quad |B_g(x_{ij})| < K[J\{H(x_{ij})\}] \quad \text{en dat} \\ J_N(1) = o(N^{\frac{1}{2}}).$$

2). Alle spesiale gevalle van die algemene toetsingsgrootheid wat in hierdie hoofstuk beskou word, voldoen aan die meer beperkende voorwaarde (3.14').

195d.

BYLAAG C.

Aan die voorwaardes

- i) $\lim_{N \rightarrow \infty} E |t'_{i.}|^{2+\delta} = O(1)$,
- ii) $\lim_{N \rightarrow \infty} E |t'_{.j}|^{2+\delta} = O(1)$ en
- iii) $\lim_{N \rightarrow \infty} E |t''_{ij}|^{2+\delta} = O(1)$ met $\delta > 0$

word voldoen onder voorwaardes (5.13), (5.14) en

$$(5.13') \quad N^{-1} \sum_p (a_p - a_{...})^4 = O(N).$$

Bewys: Die bewys dat aan voorwaardes i) en ii) voldoen word, geskied soos in bylaag A.

Deur te skrywe

$$a_{ijh} - a_{i..} - a_{.j.} + a_{...} = (a_{ijh} - a_{...}) - (a_{i..} - a_{...}) - (a_{.j.} - a_{...}),$$

kan geredelik aangetoon word dat (5.13') voldoende is vir die geldigheid van voorwaarde iii) deur op 'n soortgelyke wyse as in bylaag A te werk te gaan met gebruikmaking van die tegniek wat in die bewys van lemma 5.2 ontwikkel is.

Die geldigheid van voorwaarde (5.13') word verder implisiet aanvaar.

SUMMARY.

DISTRIBUTION-FREE TESTS FOR THE PROBLEM
OF TWO OR MORE SAMPLES.

Distribution-free tests are constructed for the k -sample problem ($k \geq 2$). Some of these tests are extensions of two-sample tests proposed by WILCOXON(1945), TERRY(1952), MOOD(1954), VAN DER WAERDEN(1957) and LEMMER and STOKER(1960).

The test statistics have asymptotic χ^2 distributions with $(k-1)$ degrees of freedom under the null hypothesis H_0 (chapter II), and also under the general hypothesis H subject to certain general conditions (chapter III - see CHERNOFF and SAVAGE(1958)). A number of these k -sample distribution-free tests for difference in location are compared with the parametric F test and in the case of tests for dispersion with the Bartlett test for homogeneity of variance by means of their asymptotic relative efficiencies. Several of the k -sample tests compare favourably with the parametric tests and are highly recommended for practical application. Necessary and sufficient conditions are given for the consistency of these tests.

Different tests are suggested for the problem of m rankings (chapter IV) and their asymptotic properties derived.

The analysis of variance in the two-factor case is studied in chapter V. Different tests for row-, column- and interaction effects are constructed, which, under H_0 , possess asymptotic χ^2 distributions and are distributed independent of each other.

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