

Can Integration of Legume Trees Increase Yield Stability in Rain-fed Maize Cropping Systems in Southern Africa?

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Abstract: Growing maize (*Zea mays*) in association with legume tree in agroforestry arrangements has been shown to increase yields in many parts of sub-Saharan Africa (SSA). However, the stability of crop yields has not been critically analyzed in the various cropping systems that integrate leguminous trees. The objective of this analysis was to compare yield stability in improved cropping systems, namely maize-Gliricidia (*Gliricidia sepium*) intercropping and fertilized monoculture maize, with the *de facto* practice of resource-poor farmers who grow maize continuously without any external input. Yield stability was determined for three long-term field trials (12-13 consecutive years) conducted at Makoka Research Station in southern Malawi and Msekera Research Station in eastern Zambia. At Makoka, the most stable yield was recorded in maize-Gliricidia intercrops. Average yield was highest in maize-Gliricidia intercropping amended with 50% of the recommended N and P fertilizer, and this was comparable with yield recorded in monoculture maize that received inorganic fertilizer. On the two sites at Msekera, the highest yield was recorded in fertilized monoculture maize followed by maize-Gliricidia intercrops. However, yields were more stable in maize-Gliricidia intercropping compared to fertilized maize on both sites at Msekera. It is concluded that maize yields remain more stable in maize-Gliricidia intercropping than in fertilized maize monoculture in the long-term although average yields may be higher with full fertilization.

Key words: Agroforestry, Cropping system, *Gliricidia sepium*, land degradation, soil fertility

INTRODUCTION

More than 95% of the farm land in sub-Saharan Africa (SSA) is rain-fed, and crop yields are generally low and variable as a consequence of variable rainfall, drought and land degradation (Wani et al., 2009). Maize accounts for over 50% of the cropped area and the calories consumed in many countries in SSA (Sileshi et al., 2010). The maize mixed farming system, which covers 10% of the land area in SSA, is currently in crisis (Dixon et al., 2001). As a result of rapid population growth, average farm sizes have fallen to under 0.5 ha in many parts of the region. The soils have also been subjected to erosion, loss of organic matter and serious fertility decline (Sanchez, 2002). Input use has fallen sharply due to increasing price. There are also signs of increasing soil acidity in some instances where there has been continuous cultivation, prolonged use of inorganic fertilizers and burning of crop residues (Munthali, 2007). Consequently maize yields have either stagnated or are in a state of decline. With climate change, maize cropping systems are also expected to experience a remarkable reduction in yield (Lobell et al., 2011). For 1°C of warming alone more than 75% of the present maize-growing areas in Africa are predicted to experience at least 20% reduction in yield under drought conditions (Lobell et al., 2011).

Land degradation can exacerbate drought as the former affects water availability, quality, and storage (Bossio et al., 2010; Diouf, 2001; Sileshi et al., 2011). Therefore, measures that mitigate land degradation are urgently needed to increase water productivity and reduce risks of crop failure under rain-fed maize cropping systems. Integration of legumes into maize cropping systems is one option for mitigating land degradation (Sileshi et al., 2011) as they add considerable amounts of organic matter and nitrogen to the soil (Akinnifesi et al., 2007; Beedy et al., 2010; Mafongoya et al., 2006; Snapp et al., 1998). Over two decades of agroforestry research in southern Africa shows that organic matter added to the system increases structural stability of the soil, resistance to rainfall impact, infiltration rates, and faunal and microbial activities (Beedy et al., 2010; Mafongoya et al., 2006; Sileshi and Mafongoya, 2006). Growing maize in association with legume trees has also been shown

to increase yields in many parts of SSA (Sileshi et al., 2008; 2010). However, the stability of yields has not been critically analyzed in cropping systems that integrate legume trees, as most of the studies are short-term in nature. Crop yield stability is an important characteristic to be considered when judging the value of a cropping system relative to others (Piepho, 1997; 1998).

In the past, the analysis of yield stability has been largely confined to trials of crop cultivars in multiple environments. However, the application of stability analyses for comparing different agronomic treatments is gaining acceptance (Guertal et al., 1994; Hildebrand, 1983; Piepho, 1998; Raman et al., 2011). The objective of the present analyses was to compare yield stability in improved cropping systems, namely maize-Gliricidia (*Gliricidia sepium*) intercropping and fertilized monoculture maize, with the *de facto* practice of resource-poor farmers. Specifically, we test the hypothesis that yields are more stable in maize-Gliricidia intercropping than in fertilized or unfertilized monoculture maize.

MATERIALS AND METHODS

Data from four long-term trials were used to test this hypothesis. The studies were located in southern Malawi and eastern Zambia (Table 1) At all sites, monoculture maize crops grown without any external inputs (control) was comparison with a fully fertilized maize crop and maize intercropped with Gliricidia (*Gliricidia sepium*), which is a nitrogen-fixing legume tree. Maize grown without any external input is the *de facto* resource-poor farmers' practice in Malawi and Zambia, and this was used as the control.

The Study Sites and Treatments

The first trial was established at Makoka Agricultural Research Station (15° 30' S, 35° 15' E; altitude 1030 m a.s.l.) in southern Malawi and run for 13 consecutive years. The rainfall is unimodal;

most of which occurs between November and April. The 30-year mean annual rainfall is 1024 mm. The cropping (wet) season extends from October to April. The soil at this site is Ferric Lixisol (FAO) with 46% sand, 46% clay and 8% silt.

This trial was established in December 1991 for studying long-term biophysical performance of *Gliricidia* intercropping on maize yield. The experimental design was a randomized complete block with three replicates. The treatments consisted of three cropping systems (fertilized maize monoculture, unfertilized maize monoculture and maize-*Gliricidia* intercropping with and without fertilizer). There were three rates of nitrogen (N) and phosphorus (P) applied as calcium ammonium nitrate (CAN) and triple superphosphate (TSP), respectively. The N rates were 0, 50% and 100% of the recommended 92 kg N ha^{-1} . The P rates were 0, 50 and 100% of recommended 40 kg P ha^{-1} for hybrid maize growing in Malawi. In this analysis, fertilized maize monoculture, maize-*Gliricidia* intercropping without fertilizer and maize-*Gliricidia* intercropping with 50% of the recommended N and P fertilizer were compared with maize grown without any input (control). *Gliricidia* seedlings were planted in pure stands at $0.9 \text{ m} \times 1.5 \text{ m}$ ($7400 \text{ trees ha}^{-1}$). The trees were cut back in September 1992 to the height of 0.30 m. Since then, the re-sprouts from the tree stumps were pruned three times during each cropping season, and leaves and twigs were incorporated in the soil. Maize hybrid NSCM 41 was planted on ridges at a spacing of 30 cm within rows and 75 cm between rows ($44,000 \text{ plants ha}^{-1}$), in both the monoculture maize as well as intercropping. The management of this trial has been described in detail in Akinnifesi et al. (2007).

The second and third trials were established at Msekera Research Station ($13^{\circ}39'S$, $32^{\circ}34' E$, altitude 1025 m) in eastern Zambia in two separate fields in 1991 and 1992. The trials are code-named Expt 91-3 and Expt 92-3 based on the year of establishment of the experiments and the length of the fallow period (3 years). Both trials had a randomized complete block design with four replicates. These trials were run for 12 consecutive years. The soils at both sites were ferric Luvisols (FAO classification) with 61% sand, 11% silt and 28% clay. The climate of the study area is humid

subtropical with three distinct seasons: the warm wet season (November to April), the cool winter (May to August) and the hot, dry season (September to October). The rainfall pattern is unimodal (averages 960 mm per year) with approximately 85% of the rains falling during December-March. *Gliricidia* seedlings were planted in pure stands at a spacing of 1 m by 1 m (10,000 trees ha⁻¹). At the end of the fallow period (36 months) trees were cut at 0.3 m height, and the leaf and twig biomass was incorporated into the soil. Since then, the re-sprouts were pruned three to four times every year and pruning biomass was incorporated into the soil. Maize (hybrid MM604) was planted on the ridges between the tree stumps every year. Continuous monoculture maize crops grown with and without fertilizer input were planted for 12 consecutive years in Expt 91-3 and Expt 92-3, respectively. Fertilized monoculture maize received the recommended rate of 200 kg ha⁻¹ yr⁻¹ compound fertiliser (N=100 g kg⁻¹, P = 90 g kg⁻¹, and K = 80 g kg⁻¹) at planting and 200 kg ha⁻¹ urea at four weeks after planting. The management of these trials have been described in detail elsewhere (Sileshi and Mafongoya 2006).

Data and Analysis

Analysis of variance was conducted on maize yield for each data set using the SAS mixed models procedure with Type 3 option (SAS Institute, 2008). In the model, the treatment (i.e. cropping system) constituted the fixed effect whereas block, year and year × treatment interactions were the random effects. Wherever the treatment main effect was significant, treatment mean yields were compared using Tukey's test at 5% level of significance.

In order to assess yield stability, various models were applied to the datasets from the three sites separately. The three datasets could not be combined because the maize varieties and experimental design used were slightly different. Several methods of stability analysis exist, but these generally fall in two broad categories, i.e. those that use either regression or variance methods (Stelluti et al., 2007). In this paper we tested the regression technique as well as three variance models.

The stability analysis by means of regression consists in relating the performance of a genotype onto an environmental index computed as the mean of all genotypes in an environment (Guertal et al., 1994; Guretzky et al., 2010). The index may be taken as a measure of the productivity of an environment or a year. In this analysis we conducted linear regression of treatment yield on the environment mean yield, calculated as the average yield of all treatments for each year (Grover et al., 2009; Guertal et al., 1994). For a valid stability analysis, changes in yield over time should not differ among the cropping systems being compared (Guertal et al., 1994). A significant treatment mean and year relationship would indicate a long-term trend with respect to the treatment and would preclude the use of stability analysis (Guertal et al., 1994; Guretzky et al., 2010).

A simple linear regression of treatment means on the environmental means for each dataset was conducted using PROC GLM of SAS (SAS Institute, 2008). The significance of regression coefficients for each treatment was determined using 5% probability level, and then tests of equality of the coefficients were performed using pre-planned comparisons. The slope of the regression line was used as a relative measure of stability, where a smaller slope was interpreted as an indication of greater yield stability (Guertal et al., 1994). The root mean square of error (RMSE) was used to judge the accuracy of estimates from the regression. RMSE is measure of the differences between values predicted by a model and the values actually observed.

Three models, the environmental variance model with unstructured variance-covariance matrix (Piepho, 1998), Shukla's stability variance model (Shukla, 1972) and the Eberhart-Russell regression model (Eberhart and Russell, 1966) were compared. The environmental variance model is the most general model of which the other two are special cases obtained by partitioning the random deviation from the mean (a_{ij}) as $u_j + e_{ij}$ for Shukla's stability variance model and as $\lambda_i u_j + d_{ij}$ for the Eberhart and Russell model. These models differ in the variance-covariance structure while the expectation structure is identical.

Following Piepho (1998) the environmental variance model can be specified as follows:

$$y_{ij} = \mu_i + a_{ij}$$

where y_{ij} ($i = 1, \dots, I; j = 1, \dots, J$) is the performance of the i^{th} treatment in the j^{th} environment, μ_i is the mean effect of the i^{th} treatment and a_{ij} is a random deviation from the mean of the i^{th} treatment in the j^{th} environment. The mean effect is the systematic component (i.e. fixed effect), for example the average yields in a cropping system in a given location. This random effect accommodates the average block effect in the j^{th} environment, the ij^{th} treatment-environment interaction effect and the experimental error term (Piepho, 1999). The a_{ij} 's for the same environment j are assumed to be correlated. The variance-covariance structure of the vector $\mathbf{a}_j = (a_{1j}, \dots, a_{ij})'$, $\text{var}(\mathbf{a}_j) = \Sigma$ is completely unstructured, i.e., the elements in the symmetric variance-covariance matrix Σ may take any value, as long as it is positive definite. The diagonal elements of Σ , σ_{ii} are environmental variances of the treatments, which may be interpreted as stability measures; the smaller the σ_{ii} , the more stable the i^{th} treatment.

Shukla's stability variance model for the treatment-environment means data can be specified as follows:

$$y_{ij} = \mu_i + u_j + e_{ij}$$

where u_j is a random environmental main effect and e_{ij} is a random residual comprising both treatment-by-environment interaction and error terms. The random effects u_j and e_{ij} are assumed to be independent with variances σ_u^2 and σ_i^2 , respectively. With this assumption, the covariance between observations in an environment is the same for all pairs of observations ($\text{cov}(y_{ij}, y_{i'j}) = \sigma_u^2$) but the variance of an observation varies among treatments as $\text{var}(y_{ij}) = \sigma_u^2 + \sigma_i^2$. The variance σ_i^2 is Shukla's (1972) stability variance and a treatment with a small stability variance is considered as stable.

Eberhart and Russell regression model (1966) for the treatment-environment mean data can be specified as follows:

$$y_{ij} = \mu_i + \lambda_i u_j + d_{ij}$$

where λ_i is a regression coefficient corresponding to the i^{th} treatment and d_{ij} is a random deviation from the regression line. The u_j is a latent environmental variable that cannot be measured directly (Piepho, 1998), but to which treatments are assumed to respond linearly, with possible differences in the strength of response indicated by differences in the sensitivity parameter λ_i . The random effects u_j and d_{ij} are assumed to be independent with variances σ_u^2 and $\sigma_{d(i)}^2$. Unlike Shukla's model, the covariance now may vary among pairs of cropping systems as $\text{cov}(y_{ij}, y_{i'j'}) = \lambda_i \lambda_{i'} \sigma_u^2$. The multiplicative term $\lambda_i u_j$ is overparameterized, therefore in the analysis the PROC MIXED identifiability constraint $\sigma_u^2 = 1$ is used. A treatment with a large absolute value of λ_i shows large sensitivity to changing environmental conditions. The variance $\sigma_{d(i)}^2$ has the interpretation of a variance of deviations from the regression line.

The models were fitted using the MIXED procedure of the SAS system (SAS Institute, 2008) following Piepho (1999). The mixed model perspective implies that trials are conducted at a random sample of environments from a target region. In all cases year was considered as a random environment. This procedure used the Restricted Maximum Likelihood (REML) method for estimating variance components.

Then Akaike's Information Criterion (AIC) was computed to assess the adequacy of the three fitted models. The model with the smallest value of the AIC is considered most desirable (Piepho, 1998). In the present study, however Akaike weights (AICw) were used as AICw has the advantage of being easy to interpret than AIC (Johnson and Omland, 2004). AICw indicates the probability that

the model is the best among the whole set of candidate models. Therefore, it provides a measure of the strength of evidence for each model (Johnson and Omland, 2004).

RESULTS

Variation in Yield with Treatment

Maize yields varied from year to year as indicated by the significant year \times treatment interactions on all sites (Table 2). The yields were also different among the treatments and years. At Makoka, the highest yield was recorded in maize-Gliricidia intercrops amended with 50% of recommended fertilizer (Table 3). Average yields in fully fertilized monoculture maize were comparable with those in maize-Gliricidia intercropping. The lowest yield was recorded in the maize grown without any external input (control), which is the *de facto* farmers' practice. At the two sites at Msekera (Expt 91-3 and Expt 91-3) the highest yield was recorded in fully fertilized monoculture maize. This was followed by maize-Gliricidia intercrops. The lowest yield was recorded in maize grown without any external input. The highest coefficient of variation (CV) was recorded in maize grown without any external input at Msekera, while the opposite was true at Makoka (Table 3). Yields declined over the years in the control and fertilized monoculture maize (data not shown).

Yield Stability

The regression technique indicated some differences among treatments over time (Table 4). Linear regression of treatment means on the environment means was highly significant ($P < 0.01$) for the trial at Makoka and Expt 91-3 at Msekera (Table 4). The slopes of the regression of mean yield on year were non-significant ($P > 0.05$) for Makoka and Expt 91-3 at Msekera. Therefore, stability analysis using the regression technique was deemed acceptable for these sites. However, for Expt 92-3 the slopes were highly significant indicating unsuitability of stability analysis using the linear

regression technique (data not shown). Therefore, results are presented only for Makoka and Expt 91-3 datasets (Table 4).

Maize-Gliricidia intercropping had the most stable yields at Makoka (Table 4). The slopes exceeded 1.0 in all treatments except maize-Gliricidia intercropping. Tests of equality of slope (Table 5) indicated significant differences only between maize-Gliricidia intercropping and fertilized maize. However, deviations from regression were smallest (indicated by RMSE) in Gliricidia + 50% N fertilizer. Variability around the intercept and slope (indicated by the standard errors) was smaller for Gliricidia + 50% fertilizer than that for the Gliricidia, fertilized maize and the control (Table 4). Examination of the slopes (Table 4) indicated that maize grown without any external inputs had the most stable yields in Expt 91-3 at Msekera. Variability around the intercept and slope was largest in the fertilizer monoculture maize. Deviations from regression were also larger in the fertilizer monoculture maize than in maize-Gliricidia intercropping and the control (Table 4). Tests of equality of slope (Table 5) indicated no significant differences in all comparisons in Expt 91-3.

For a given variety of maize, the environmental variance was larger when fertilizer was applied (e.g. fertilized maize and Gliricidia + 50% fertilizer) suggesting that these treatments were not as stable as the maize-Gliricidia intercropping (Table 6). According to the environmental variance model, fully fertilized monoculture maize crop is the least stable in all three trials. Maize-Gliricidia intercrop without fertilizer was the most stable at Makoka. Maize crop associated with Gliricidia also had a relatively small regression coefficient (λ_i) which indicates least sensitive to changing environmental conditions. The environmental variance and Shukla's stability variance models indicated the highest stability in the control, followed by the maize-Gliricidia intercropping in Expt 91-3 and Expt 92-3 (Table 6). However, the Eberhart and Russell model indicated highest stability in maize-Gliricidia intercropping. In the case of control, parameter estimates had large standard errors indicating that 12-13 environments are not an adequate sample to obtain reliable stability estimates using Shukla's stability variance model.

According to the model comparisons based on AICw, the best fits were obtained by the environmental variance model for the Makoka site and Eberhart and Russell model for Expt 91-3 and Expt 92-3. The AICw shows that the environmental variance model has a 57% chance of being the appropriate model for the Makoka dataset relative to a 24% and 19% chance for the stability variance and Eberhart and Russell models, respectively (Table 7). The Eberhart and Russell model had 53 and 63% chance of being the appropriate model for Expt 91-3 and Expt 92-3, respectively (Table 7). Therefore, the Eberhart and Russell stability model was considered the most appropriate for interpreting yield stability on both sites at Msekera. Accordingly, yields were judged to be more stable in maize-Gliricidia intercropping compared to the fertilized monoculture maize at Msekera.

The generalized least squares estimates of μ_i for the various treatments in the three trials are presented in Table 8. While the estimates of μ_i are consistent across the different models, there are differences in standard errors of estimates between models. The standard errors of estimates from the stability variance model were different from those of the Eberhart and Russell and the environmental variance model (Table 8). Based on the model selected using AICw (Table 7), the standard errors provided by the environmental variance model are the most appropriate for the Makoka dataset (Table 8). In the case of Expt 91-3 and Expt 92-3, the most appropriate standard errors are those provided by the Eberhart and Russell model.

DISCUSSION

The results revealed that intercropping maize with Gliricida can significantly increase maize yield over maize grown without any external inputs in Malawi and Zambia. This may be attributed to the enhancement of nutrient (Akinnifesi et al., 2007; Sileshi and Mafongoya, 2006) and water availability (Chirwa et al. 2007) in the maize-Gliricidia intercrop. The re-sprout biomass (green leaf + twigs) from Gliricidia incorporated in the soil was estimated to add up to 302 kg ha⁻¹ N yr⁻¹ in the trial at Makoka (Akinnifesi et al., 2006) and 124 kg ha⁻¹ N yr⁻¹ in Expt 92-3 at Msekera (Sileshi and

Mafongoya, 2006). Similarly, the re-sprout biomass was estimated to add up to 21 kg P ha⁻¹ yr⁻¹ at Makoka (Akinnifesi et al., 2006) and up to 8 kg P ha⁻¹ yr⁻¹ in Expt 92-3 (Sileshi and Mafongoya, 2006). The pruning biomass has also been shown to increase organic matter (SOM) and faunal activity in the soil (Beedy et al. 2010; Sileshi and Mafongoya, 2006). For example, SOM, particulate organic matter (POM), POM-C and POM-N were 12, 40, 62 and 86% higher in the maize-Gliricidia intercrop compared to monoculture maize at the Makoka site (Beedy et al. 2010). These increases in SOM and faunal activity in turn may lead to improvements in soil structure and water dynamics. For example, infiltration rates and water use efficiency were higher in the maize-Gliricidia intercrops than monoculture maize at the study sites (Chirwa et al. 2007; Sileshi and Mafongoya, 2006).

Although yields were high in fertilized maize they were unstable in all three trials. This highlights the fact that high yield is not necessarily an indicator of sustained productivity as it may be associated with low stability. At the Msekera site, yields in the control and fertilized maize showed declining trend over the years (Sileshi and Mafongoya, 2006). This is in agreement with studies conducted elsewhere. For example, continuous maize cropping of maize with inorganic fertilizer at Ibadan resulted in significant yield decline over a period of 16 years (Vanlauwe et al., 2005). Based on a 14 year study in Punjab, Bhandari et al. (2002) noted decline in rice yield even when the recommended rates of N, P, and K were applied. This was attributed to loss of total soil N and organic matter (Bhandari et al., 2002). Application of fertilizer without addition of organic matter may not be sustainable because only organic matter adds carbon, feeds soil biota and helps to retain soil moisture (Sileshi and Mafongoya, 2006; Sileshi et al., 2011).

The evaluation of yield stability requires use of appropriate measures of variability around a mean. The coefficient of variation (CV) is of limited use as it gives only a relatively simple expression of the variability around a mean yield (Rao and Willey, 1980). The CV could also give a misleading impression as it does not take the covariance structure into account. The concept of stability implies a random (i.e. unpredictable) element in the performance of a cropping system.

Therefore, we used variance components as they measure variability across environments more efficiently than the CV (Table 3) and the regression technique (Table 4). The mixed modeling framework allows accurate estimation of this random component. While various procedures of stability analysis may be applicable, the results may not always be concordant (Stelluti et al., 2007). One has to use appropriate variance-covariance matrix for valid inferences (Piepho, 1999). The results of the present analyses suggest that the environmental variance model was more suitable for the Makoka trial than the other models. On the other hand, the Eberhart and Russell model was the most suitable for Expt 91-3 and Expt 92-3 datasets. This highlights the need for selecting models appropriate for the data at hand for better interpretation. According to the mixed modeling framework the problem of choosing an appropriate stability measure can be regarded as the problem of identifying the most appropriate variance-covariance structure and this choice is data-dependent. Therefore, the usefulness of any measure of stability depends crucially on how well the underlying model approximates the real data (Piepho, 1998). However, this difference does not always happen. The major advantage of mixed model approach is its applicability to unbalanced data for stability measures by using restricted maximum likelihood (REML) method. The data may be unbalanced owing to missing observations on some plots, varying number of replications among trials and some cropping system by environment combinations that may not be tested.

To our knowledge this is the first study of its kind that analyzed long-term trends in crop yield stability in cereal-legume tree associations in southern Africa. In fact most of the studies assessing cereal-legume tree associations are short term (see Sileshi et al., 2008). As highlighted by a review of intercropping studies (Connolly et al., 2001), concerns about stability and sustainability are not as central in many of the intercropping studies as might have been expected. Short-term studies do not allow one to identify reliable cropping systems, i.e. those that combine high levels of mean yield and yield stability. It is also not possible to determine with any degree of certainty trends in soil processes or weather changes with short-term studies (Girma et al., 2007). The large standard errors

of the variance estimates observed in most of the treatments indicate that 12-13 environments are not an adequate sample to obtain reliable stability estimates. This emphasizes the need for long-term monitoring of maize yield in the cropping systems investigated. The value of long-term monitoring in revealing dynamic soil processes has been demonstrated (Davis et al., 2003; Edmeades, 2003; Girma et al., 2007; Miles and Brown, 2011) in studies ranging from 20 to 120 years in the USA and Europe. Such long-term studies are virtually non-existent in SSA.

Understanding temporal variability in crop yields has implications for sustainable crop production and food security in Africa. Long-term trials are vital for identifying cropping systems with high and stable crop yields and low production risk (Grover et al., 2009). Such experiments now have additional and immediate relevance for understanding and predicting the consequences of global change taking place in Africa. A network of long-term trials may also provide useful information on the vulnerability of staple crops and current management practices to future changes. The value of such trials has recently been demonstrated by Lobell et al. (2011) on the effects of potential warming on African maize. Therefore, we recommend establishment of well-designed long-term trials that allow stability analyses to help in assessing current and expected vulnerabilities of cropping systems with a changing climate in Africa. Such information can help in exploring possible technological alternatives and policy interventions to improve the adaptability and sustainability of cropping systems.

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Table 1. Baseline soil properties (top 20 cm) of the study sites at Makoka Agricultural Research Station in Malawi and Msekera Research Station in Eastern Zambia

Soil properties	Makoka site	Msekera sites
Soil type (FAO)	Ferric Lixisols	Ferric Luvisols
Sand (%)	46	61
Clay (%)	46	28
Organic C (g kg ⁻¹)	8.8	10.2
pH (H ₂ O)	5.9	5.3
Mg (c mol _c /kg)	1.6	1.7
Ca (c mol _c /kg)	4.4	3.0
K (c mol _c /kg)	0.3	1.5

Table 2. Type 3 analysis of variance for maize yield in cropping systems (Treatment) at Makoka Agricultural Research Station in Malawi, and in Expt 91-3 and Expt 92-3 at Msekera Research Station in Eastern Zambia

Source of Variation	Makoka		Expt 91-3		Expt 92-3	
	DF	MS	DF	MS	DF	MS
Block	2	4.09**	3	0.45 ns	3	4.86***
Year	12	26.88***	11	3.95**	11	8.22***
Treatment	3	27.82***	2	107.34***	2	118.53***
Year × Treatment	36	2.65***	22	0.95**	22	1.71***
Residual	102	0.743	93	0.46	105	0.48

DF = degrees of freedom; MS = mean square

** Significant at the 0.01 probability level.

*** Significant at the 0.001 probability level.

ns = Non significant

Table 3. Mean maize yields (Mg ha⁻¹) in various treatments at Makoka Agricultural Research Station in Malawi, and Expt 91-3 and Expt 92-3 at Msekera Research Station in Eastern Zambia

Treatment [#]	Makoka	Expt 91-3	Expt 92-3
Control	2.62 (29.0) a ⁺	1.20 (66.7) a ⁺	0.98 (56.1) a ⁺
Fertilizer	4.17 (34.8) bc	4.42 (31.4) c	4.08 (40.7) c
Gliricidia	3.72 (41.9) b	3.06 (46.7) b	2.95 (40.7) b
Gliricidia + 50%F	4.58 (32.5) c	NA	NA

[#]Treatments: Control = unfertilized continuous monoculture maize; Gliricidia = maize-Gliricidia intercropping without fertilizer; Fertilizer = fertilized continuous monoculture maize; Gliricidia + 50%F = maize-Gliricidia intercropping + 50% of the recommended fertilizer

⁺ Figures in parentheses are coefficients of variation. These were included upon request by the editor.

Means followed by the same letter in a column are not significantly different according to Tukey's test at 0.05 probability level.

NA = not applicable

Table 4. Linear regressions for treatment mean maize yield on the environmental mean yields at Makoka Agricultural Research Station in Malawi, and Expt 91-3 and Expt 92-3 at Msekera Research Station in Eastern Zambia

Experiment	Treatment [#]	Intercept	SE	Slope	SE	R ²	RMSE	$P > t $ [*]
Makoka	Gliricidia	1.189	0.808	0.671	0.200	0.505	1.039	0.0065
	Fertilizer	-0.441	0.586	1.222	0.145	0.866	0.753	<0.0001
	Gliricidia+50%F	0.480	0.380	1.087	0.094	0.924	0.489	<0.0001
	Control	-1.214	0.604	1.016	0.150	0.807	0.777	<0.0001
Expt 91-3	Gliricidia	-0.402	0.483	1.197	0.164	0.856	0.308	<0.0001
	Fertilizer	1.087	0.725	1.151	0.246	0.709	0.462	0.0011
	Control	-0.699	0.568	0.657	0.192	0.564	0.362	0.0077

[#]Treatments: Control = unfertilized continuous monoculture maize, Gliricidia = maize-Gliricidia intercropping without fertilizer; Fertilizer = fertilized continuous monoculture maize; Gliricidia + 50%F = maize-Gliricidia intercropping + 50% of the recommended fertilizer

RMSE = root mean square of error

^{*} $P > |t|$ probability of a greater absolute value of t for slope.

Table 5. Tests of equality of slopes among treatments at Makoka Agricultural Research Station in Malawi, and Expt 91-3 at Msekera Research Station in Eastern Zambia

Experiment	Comparison [#]	$P > F$ [*]
Makoka	Gliricidia vs. Gliricidia + 50%F	0.0597
	Gliricidia + 50%F vs. Fertilizer	0.5328
	Fertilizer vs. Control	0.3439
	Gliricidia vs. Fertilizer	0.0139
	Gliricidia vs. Control	0.1156
	Gliricidia + 50%F vs. Control	0.7444
	Overall	0.0822
Expt 91-3	Gliricidia vs. Fertilizer	0.8731
	Fertilizer vs. Control	0.0974
	Gliricidia vs. Control	0.0712
	Overall	0.1340

[#] Overall, equality of slopes among cropping systems; Control = unfertilized maize.

^{*} $P > F$, probability of a greater F statistic.

Table 6. REML parameter estimates (standard errors in parenthesis) of the variance-covariance structure of different stability models for the datasets from Makoka Agricultural Research Station in Malawi, and Expt 91-3 and Expt 91-3 at Msekera Research Station in Eastern Zambia

Experiment	Treatment [#]	Environmental variance, σ_{ii}	Shukla's stability variance (σ_i^2)	Eberhart-Russell (σ_i^2)	Eberhart-Russell (λ_i)
Makoka	Control	2.87 (1.17)	0.81 (0.46)	0.67 (0.36)	1.49 (0.39)
	Fertilizer	3.87 (1.58)	0.69 (0.36)	0.40 (0.39)	1.86 (0.43)
	Gliricidia	2.00 (0.82)	1.58 (0.72)	1.32 (0.60)	0.83 (0.40)
	Gliricidia+50%F	2.87 (1.17)	0.23 (0.24)	0.35 (0.31)	1.59 (0.38)
Expt 91-3	Control	0.27 (0.12)	0.09 (0.09)	0.14 (0.06)	0.36 (0.15)
	Fertilizer	0.66 (0.30)	0.45 (0.25)	0.38 (0.17)	0.53 (0.23)
	Gliricidia	0.59 (0.27)	0.23 (0.14)	0	0.77 (0.17)
Expt 92-3	Control	0.32 (0.135)	0.02 (0.084)	0.14 (0.061)	0.42 (0.145)
	Fertilizer	1.84 (0.785)	1.30 (0.597)	0.86 (0.366)	0.99 (0.350)
	Gliricidia	0.76 (0.322)	0.33 (0.171)	0	0.87 (0.185)

[#]Treatments: Control = unfertilized continuous monoculture maize, Gliricidia = maize-Gliricidia intercropping without fertilizer; Fertilizer = fertilized continuous monoculture maize; Gliricidia + 50%F = maize-Gliricidia intercropping + 50% of the recommended fertilizer

A small variance (σ_i^2 value) indicates high relative stability.

Table 7. Akaike's Information Criterion (AIC) and Akaike weights (AICw) for the three models fitted to the datasets from Makoka in Malawi, and Expt 91-3 and Expt 92-3 at Msekera Research Station in Eastern Zambia

Experiment	Model	AIC	AICw
Makoka	Environmental variance	171.4	0.568
	Stability variance	173.1	0.243
	Eberhart and Russell	173.6	0.189
Expt 91-3	Environmental variance	69.8	0.216
	Stability variance	69.5	0.251
	Eberhart and Russell	68.0	0.532
Expt 92-3	Environmental variance	86.9	0.242
	Stability variance	88.1	0.133
	Eberhart and Russell	85.0	0.625

Table 8. REML estimates of maize yields (Mg ha^{-1}) and standard errors (in parenthesis) for different stability models of the datasets from Makoka Agricultural Research Station in Malawi, and Expt 91-3 and Expt 91-3 at Msekera Research Station in Eastern Zambia

Experiment	Treatment	Environmental variance μ_i	Stability variance μ_i	Eberhart-Russell μ_i
Makoka	Control	2.62 (0.470)	2.62 (0.502)	2.62 (0.470)
	Fertilizer	4.17 (0.546)	4.17 (0.493)	4.17 (0.546)
	Gliricidia	3.72 (0.392)	3.72 (0.558)	3.72 (0.392)
	Gliricidia + 50%F	4.58 (0.470)	4.58 (0.455)	4.58 (0.470)
Expt 91-3	Control	1.20 (0.157)	1.20 (0.170)	1.20 (0.157)
	Fertilizer	4.42 (0.245)	4.42 (0.248)	4.42 (0.245)
	Gliricidia	3.07 (0.232)	3.07 (0.205)	3.07 (0.232)
Expt 92-3	Control	0.98 (0.163)	0.98 (0.165)	0.98 (0.163)
	Fertilizer	4.08 (0.392)	4.08 (0.367)	4.08 (0.392)
	Gliricidia	2.95 (0.251)	2.95 (0.229)	2.95 (0.251)