

Bibliography

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Appendix A

Probability distributions

A.1 The normal distribution

The p.d.f. of a k -dimensional multivariate normal distribution is specified as follows [27]:

$$g(x|\mu, \Sigma) = (2\pi)^{-D/2} |\Sigma|^{-1/2} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1} (x-\mu)}. \quad (\text{A.1})$$

A normal distribution with mean μ and covariance matrix Σ is usually written as $\mathcal{N}(\mu, \Sigma)$. If we assume that the individual elements of the random variable x are independent (diagonal covariance matrix), we can write Equation A.1 as follows

$$g(x|\mu, \Sigma) = (2\pi)^{-D/2} \left(\prod_{d=1}^D \sigma_d \right)^{-1} e^{-\frac{1}{2} \sum_{d=1}^D \sigma_d^{-2} (x_d - \mu_d)^2} \quad (\text{A.2})$$

and

$$\log(g(x|\mu, \Sigma)) = -\frac{D}{2}\log(2\pi) - \frac{1}{2}\log\left(\prod_{d=1}^D \sigma_d^2\right) - \frac{1}{2}\sum_{d=1}^D \sigma_d^{-2}(x_d - \mu_d)^2 \quad (\text{A.3})$$

A.2 The Wishart distribution

Given a random sample of k -dimensional random vectors (X_1, X_2, \dots, X_n) from a multivariate normal distribution with zero mean and covariance matrix Σ , the random variable V has a Wishart distribution [27] with n degrees of freedom and parametric matrix Σ when,

$$V = \sum_{i=1}^n X_i X_i^T. \quad (\text{A.4})$$

For any $k \times k$, positive definite, symmetric matrix v , we have

$$g(v|n, \Sigma) = c|\Sigma|^{-n/2}|v|^{(n-D-1)/2}e^{-\frac{1}{2}\text{tr}(\Sigma^{-1}v)}. \quad (\text{A.5})$$

Here, $\text{tr}(\Sigma^{-1}v)$ is the trace of the matrix $\Sigma^{-1}v$. The value c is a normalizing constant which ensures that the integral of $g(v|n, \Sigma)$ is equal to one.

If Σ and v are assumed to be diagonal, then Equation A.5 can be rewritten as follows:

$$g(v|n, \Sigma) = c\left(\prod_{d=1}^D \tau_d\right)^{n/2}\left(\prod_{d=1}^D v_d\right)^{(n-D-1)/2}e^{-\frac{1}{2}\sum_{d=1}^D v_d\tau_d} \quad (\text{A.6})$$

where τ is the precision matrix for the Wishart distribution ($\tau = \Sigma^{-1}$) and

$$\log(g(v|n, \Sigma)) = \log(c) + \frac{n}{2}\log\left(\prod_{d=1}^D \tau_d\right) + \frac{n-D-1}{2}\log\left(\prod_{d=1}^D v_d\right) - \frac{1}{2}\sum_{d=1}^D v_d\tau_d. \quad (\text{A.7})$$

A.3 Dirichlet distribution

Given the random vector $X = (X_1, X_2, \dots, X_k)^T$ with the following properties: For a point $x = (x_1, x_2, \dots, x_k)^T$ in \mathfrak{R}^k , $x_i > 0; i = 1, \dots, k$ and $\sum_{i=1}^k x_i = 1$, then the random vector X has a Dirichlet distribution [27]:

$$g(x|\alpha) = \frac{\Gamma(\alpha_1 + \alpha_2 + \dots + \alpha_k)}{\Gamma(\alpha_1)\Gamma(\alpha_2) \dots \Gamma(\alpha_k)} x_1^{\alpha_1-1} x_2^{\alpha_2-1} \dots x_k^{\alpha_k-1}, \quad (\text{A.8})$$

where $\Gamma(\alpha)$ is the gamma function and α is the parametric vector of the distribution and

$$\begin{aligned} \log(g(x|\alpha)) = & \log(\Gamma(\alpha_1 + \alpha_2 + \dots + \alpha_k)) - \log(\Gamma(\alpha_1)) - \dots - \log(\Gamma(\alpha_k)) \\ & + (\alpha_1 - 1)\log(x_1) + \dots + (\alpha_k - 1)\log(x_k). \end{aligned} \quad (\text{A.9})$$

A.4 The gamma distribution

A random variable X has a gamma distribution [27] with parameters α and β ($\alpha > 0, \beta > 0$) if X has an absolutely continuous distribution whose p.d.f. is

$$g(x|\alpha, \beta) = \begin{cases} \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases} \quad (\text{A.10})$$

where $\Gamma(\alpha)$ is the gamma function, which is defined as

$$\Gamma(z) = \int_0^\infty t^{z-1} e^{-t} dt. \quad (\text{A.11})$$

If a random variable X has a gamma distribution as given in (A.10), then

$$E(X) = \frac{\alpha}{\beta} \quad (\text{A.12})$$

$$\text{Var}(X) = \frac{\alpha}{\beta^2}. \quad (\text{A.13})$$

If the mean μ and variance σ^2 of a gamma distribution are known, the distribution parameters α and β can easily be obtained as follows:

$$\alpha = \frac{\mu^2}{\sigma^2} \quad (\text{A.14})$$

$$\beta = \frac{\mu}{\sigma^2}. \quad (\text{A.15})$$

In this thesis, a gamma distribution with parameters α and β has been referred to as $\mathcal{G}(\alpha, \beta)$.

A.5 Conjugate families of distributions

If the prior distribution of θ belongs to a conjugate family of distributions [27], then for any sample size n and any values of the observations in the sample, the posterior distribution of θ must also belong to the same family. A family of distributions with this property is said to be closed under sampling.